### Linear Triangular Transport at Scale

Berent Lunde



EnKF Workshop, Os Norway June 17, 2024





EnKF's and convergence

Method scalability

KLD, Structure, & EnIF

Innovations for scalability

Synthetic reservoir application: Sequential EnIF



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No.



## How come they do not converge?

• Consider the stochastic heat equation,

$$du_t(\mathbf{x}) = \alpha \operatorname{div} \nabla u_t(\mathbf{x}) \, dt + \sigma \, dW_t. \tag{1}$$

Let u be a p-vector of values indexed by time t and space x.



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• Let  $\boldsymbol{U}_n$  be an  $n \times p$  matrix of *n*-samples,  $\boldsymbol{u}^{(i)}$ . The sample covariance

$$\hat{\boldsymbol{\Sigma}}_{\boldsymbol{u}}^{*} = \frac{1}{n-1} \sum_{i=1}^{n} (\boldsymbol{u}^{(i)} - \overline{\boldsymbol{u}}) (\boldsymbol{u}^{(i)} - \overline{\boldsymbol{u}})^{\top}, \qquad (2)$$

does not *trivially* converge to the population covariance  $\Sigma_{\boldsymbol{u}} = E[(\boldsymbol{u} - E[\boldsymbol{u}])(\boldsymbol{u} - E[\boldsymbol{u}])^{\top}]$  when both *n* and  $p \to \infty$ .



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• EnKF's (estimated) Kalman gains are a function of sample covariance++.



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- Often, numerical integration promised to work as  $\Delta \mathbf{x} \rightarrow 0$ .
- But  $\Delta \mathbf{x} \to 0$  implies  $p \to \infty$  in a statistical setting.
- For ensemble based methods, we need to guarantee convergence under simultaneous limits  $p \to \infty$  and  $n \to \infty$ .



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- 2. Overconfidence (loss of variability) due to belief in (random) connections and propagating a Bayesian update through them.



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From neglecting effects of  $p \to \infty$  we have incurred a very real problem.



## Spurious correlations and ensemble collapse





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## Scalability also means good statistical properties!

A KLD / Likelihood perspective:

• Minimize

$$D_{KL}(P \parallel Q) = \int \int p(\mathbf{u}, \mathbf{y}) \log \left( rac{p(\mathbf{u}, \mathbf{y})}{q(\mathbf{u}, \mathbf{y})} 
ight) d\mathbf{u} d\mathbf{y},$$

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- When Q is Gaussian, we have a Kalman-type method.
- *θ* is typically estimated. **Statistical convergence** matters (not just asymptotic expectations!).



### Ensemble Smoother (ES)





## Berent's biased map of methods



### Adaptive Localization n = 100





## Berent's biased map of methods





Presented a solution that solves non-linearity and non-Gaussianity!







How does it scale computationally?







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Must search for the rearrangement.





#### How does it scale computationally?

- Must search for the rearrangement.
- Must learn the degree of non-linearity.



## Berent's biased map of methods



$$\boldsymbol{u}_{\text{posterior}}^{i} = \boldsymbol{u}_{\text{prior}}^{i} + \boldsymbol{K}_{\text{EnKF}}(\boldsymbol{d}^{i} - \boldsymbol{h}(\boldsymbol{u}_{\text{prior}}^{i}))$$

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$$\mathbf{K}_{EnKF} = \mathbf{U}\mathbf{Y}^{ op} \left(\mathbf{Y}\mathbf{Y}^{ op} + \Sigma_{\epsilon}
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★ Regularised linear regression on map  $h^{-1}$  :  $d \mapsto u$  (Lasso).



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 $\star\star\star$  Information theoretic triangular measure transport (IT-TMT).



### EnlF:Cholesky & EnlF::Direct n = 100











## Information Theoretic (adaptive) Triangular Measure Transport



- The statistical convergence of methods cannot be neglected. Both *p* and *n* must be considered. Methods are spatio-temporal.
- When only considering asymptotic expectations, everything seems to be okay. Do not forget variance of statistics.
- The map of methods is my subjective and biased view of things.



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Model complexity

References: Akaike 1974; Takeuchi 1976; Claeskens and Hjort 2008; Hastie, Tibshirani, eqeinol: 2009

#### Information criteria and tools

- Given model complexity
- Reason about test loss

• TIC: 
$$\operatorname{tr}(\nabla_{\boldsymbol{\theta}}^2/\operatorname{cov}(\hat{\boldsymbol{\theta}}))$$
  
• AIC:  $\boldsymbol{\rho} = \operatorname{len}(\boldsymbol{\theta})$ 

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equinor. 2009

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- $\rightarrow$  Reparametrisation, with smaller p.
- $\rightarrow\,$  Regularization: Trade bias for variance

Let L be a differential operator, then the solution to

 $Lu(x) = W(\cdot)$ 

is a Gaussian random field and it has the Markov property.

• Heuristically: derivatives (local) create the Markov properties (local).



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- SPDE approach: approximate non-Markov field (solutions) by Markov fields.
- Computationally important:  $\Lambda_u = \Sigma_u^{-1}$  is **sparse** for GMRF.
- If  $\Lambda_u$  is sparse, then  $p = len(\theta)$  is much smaller than in the covariance parametrisation. Training bias in KLD is positively monotone in p.

A stochastic wave equation

$$d^2 u_t(\mathbf{x}) = c \operatorname{div} \nabla u_t(\mathbf{x}) dt^2 + \sigma dW_t$$

suggests a simple finite difference discretization, in the 1-d case:

$$u_{i}^{j+1} = 2u_{i}^{j} - u_{i}^{j-1} + \frac{c^{2}\Delta t^{2}}{\Delta x^{2}}(u_{i+1}^{j} - 2u_{i}^{j} + u_{i-1}^{j}) + \sigma\sqrt{dt}Z, \ Z \sim \mathcal{N}(0, 1)$$

So  $u_i^{j+1}$  is only a function of  $u_i^j$ ,  $u_i^{j-1}$ ,  $u_{i-1}^j$ , and  $u_{i+1}^j$ . Not all of u.



## Discretization incur spatio-temporal conditional independence





## What are we doing with EnKFs?





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- So there is no *local* solution, every displacement is a function of the global state.
- We also consider teleportation of information.
- We try to learn the physics from scratch. We need a lot of data.



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- **Smoothing**: *natural* graph from (s)pde.
- Filtering: A complete graph! No exact conditional independence.
- Parameter estimation: Sampling from independence or variograms.
- We only require a (parsimonious!) approximation.



Let **u** and **y** be jointly Gaussian. Then, a sample  $(u_i, y_i)$  is mapped to a sample from the conditional p(u|y), having observed  $y^*$ , via the formula

$$oldsymbol{u}_i + oldsymbol{K}(oldsymbol{y}^* - oldsymbol{y}_i) \sim oldsymbol{
ho}(oldsymbol{u}|oldsymbol{y}^*).$$

where the "Kalman gain" **K** is defined as  $\mathbf{K} = \Sigma_{uy} \Sigma_y^{-1}$ , which is estimated.



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- $\Lambda_{t|t} = \Lambda_{t|t-1} + H^{\top} \Lambda_{r_t} H$  is dense if H is dense.
  - Havard Rue and Held 2005 *H* is dense for geostatistcs. This won't work.
  - IF equations  $\pmb{u}_{t|t}^{(i)} = \Lambda_{t|t}^{-1} \pmb{\eta}_{t|t}^{(i)}$  computationally infeasible.



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- But *H* is estimated. Remember KLD. Choose regularisation to obtain sparse *H*.

## The Ensemble Information Filter

Sample from belief  $oldsymbol{u}_{t-1|t-1}^{(i)} \sim 
ho(oldsymbol{u}_{t-1|t-1}) \, i=1,\ldots,n$ 

 $\begin{array}{c} \textbf{Predict} \\ \textbf{u}_{t|t-1}^{(i)} = g(\textbf{u}_{t-1|t-1}^{(i)}) \end{array}$ 

Using sample  $\{u_{t|t-1}^{(i)}\}_{i=1}^{n}$  estimate  $\hat{\Lambda}_{t|t-1}$  w.r.t. graph  $\mathcal{G}$ And  $\hat{H}$  as a sparse linear map

$$\begin{split} & \textbf{Update realizations and precision} \\ & \boldsymbol{\eta}_{t|t-1}^{(l)} = \hat{\boldsymbol{\Lambda}}_{t|t-1} \boldsymbol{u}_{t|t-1}^{(l)} \\ & \hat{\boldsymbol{\eta}}_{t|t} = \hat{\boldsymbol{\eta}}_{t|t-1} + \hat{\boldsymbol{H}}^\top \boldsymbol{\Lambda} \boldsymbol{r}(\boldsymbol{y}_t - \boldsymbol{r}^{(l)}) \\ & \hat{\boldsymbol{\Lambda}}_{t|t} = \hat{\boldsymbol{\Lambda}}_{t|t-1} + \hat{\boldsymbol{H}}^\top \boldsymbol{\Lambda} \boldsymbol{r} \hat{\boldsymbol{H}} \end{split}$$

Bring realizations back to original space

 $\boldsymbol{u}_{t|t}^{(i)} = \hat{\boldsymbol{\Lambda}}_{t|t}^{-1} \boldsymbol{\eta}_{t|t}^{(i)}$ 


• KLD warrants the use of structure and regularisation.

- Structure can come from the model, e.g. (S)PDE.
- Derivatives (local) leads to Markov properties (local), perhaps approximately.
- EnIF is a reparametrisation of the Gaussian update in EnKF. Regularised and encoding (Markov) structure.
- Sparsity is a necessity for computation.



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# An Equinor history matching problem

- About 10 million parameters p.
  - About 100-1000 static parameters
  - $\bullet~$  Some surfaces of size about  $300\times300$
  - Some 3D fields of size about  $100 \times 100 \times 100$
- Ensemble size *n* about 100-200.
- Number of responses *m* about 100-1000, more if seismic is included.

And how to **understand the update** for a domain expert?



### EnIF using off-the-shelf libraries vs. with innovations



First thought: L1/LASSO regression Tibshirani 1996

• For sparsity the go-to solution.

• It is efficient, but...



# First hurdle: learning H

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# First thought: L1/LASSO regression Tibshirani 1996

- For sparsity the go-to solution.
- It is efficient, but...
- Not that efficient.



Algorithm Boosting Monotone-LASSO

- 1: Initialize  $\hat{oldsymbol{eta}}_0 = oldsymbol{0}$
- 2: while  $ext{mse}_{\mathsf{cv}-n}(\pmb{X},\pmb{y};\hat{\pmb{eta}}_k) > ext{mse}_{\mathsf{cv}-n}(\pmb{X},\pmb{y};\hat{\pmb{eta}}_{k+1})$  do
- 3: Calculate all 1d linear regressions
- 4: Select  $\beta_j$  as the one reducing training mse the most
- 5:  $\hat{\boldsymbol{\beta}}_{k+1,j} + = \epsilon \beta_j$
- 6: end while
- 7: return  $\hat{oldsymbol{eta}}$

LASSO, LARS, FS- $\epsilon$  and Boosting relations. Hastie, Taylor, et al. 2007



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•  $\hat{\theta}_{-i} \rightarrow_n \hat{\theta} - n^{-1} IF(y_i, \mathbf{x}_i)$ , where the influence *IF* is found using the asymptotic properties of  $\hat{\beta}_j$  as an M-estimators. cv-n and TIC relation through IF. Claeskens and Hjort 2008

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### Information theoretic stopping criterion



#### Information criteria and tools

- Boost out monotone-LASSO solution paths.
- With information theoretic stopping criterion!



# Second hurdle: Graph optimisation and fill-in



Given a permutation optimised for a sparse Cholesky factoer

 $\boldsymbol{L}_* \boldsymbol{L}_*^ op = \boldsymbol{P}_*^ op \boldsymbol{\Lambda}_{\boldsymbol{u}} \boldsymbol{P}_*$ 

We can find a relation to the linear triangular transport map  $\pmb{C}(\pi^*)$ 

 $\boldsymbol{\Lambda}_{\boldsymbol{u}} = \boldsymbol{P}_{\boldsymbol{r}} \boldsymbol{P}_{\ast} \boldsymbol{C}(\pi^{\ast})^{\top} \boldsymbol{C}(\pi^{\ast}) \boldsymbol{P}_{\ast}^{\top} \boldsymbol{P}_{\boldsymbol{r}}.$ 

where  $P_r$  is the reverse permutation matrix.

Learn  $C(\pi^*)$  row-by-row like in TMT, with the same sparsity as  $L_*$  (but reversed and transposed).



# Fill-in reducing optimised permutation

#### Fill-in reducing algorithms

• Finding optimal permutation is NP-hard.



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  - The dense Cholesky would have  $320 \times 10^9$ .
- This is a limitation also for triangular measure transport.



# Optimise $\Lambda$ directly column-by-column



Searched far and wide for an algorithm estimation  $\Lambda$  conditioned on  $\mathcal{G}.$ 

Surprisingly little literature, and what exist often does not scale (e.g. ESL Alg. 17.1, the basis of much more well known Graphical-Lasso Alg. 17.2. Hastie, Tibshirani, et al. 2009

Reverted to GraphSPME library. **Benefits:** very fast and no fill-in. **disadvantages:** Does not optimise the likelihood directly, symmetry, and condition number.



The final step of EnIF is to map from "canonical" realisations to physical ones.

$$oldsymbol{u}_{t|t} = oldsymbol{\Lambda}_{t|t}^{-1} oldsymbol{
u}_{t|t}$$

The natural solver is the (permutation optimised) sparse Cholesky solver.



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An **iterative solver** is the solution:  $\Lambda_{t|t}$  is SPD and sparse, thus **Conjugate gradient.** 



A "localization" effect from assuming Markov properties:

• Covariance effect through path (think AR-p) exhibits exponential decay in steps.



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The system of equations  $\Lambda {m u}=\eta$  in block-form:

$$\begin{bmatrix} \boldsymbol{\Lambda}_{11} & \boldsymbol{\Lambda}_{12} \\ \boldsymbol{\Lambda}_{21} & \boldsymbol{\Lambda}_{22} \end{bmatrix} \begin{bmatrix} \boldsymbol{u}_1 \\ \boldsymbol{u}_2 \end{bmatrix} = \begin{bmatrix} \boldsymbol{\eta}_1 \\ \boldsymbol{\eta}_2 \end{bmatrix}$$

Given  $\boldsymbol{u}_1$  is known from a previous computation, we can update  $\boldsymbol{u}_2$  as follows:

$$oldsymbol{u}_2 = oldsymbol{\Lambda}_{22}^{-1} \left(oldsymbol{\eta}_2 - oldsymbol{\Lambda}_{21}oldsymbol{u}_1
ight)$$

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Potentially a much smaller system.

EnKF's and convergence

Method scalability

KLD, Structure, & EnIF

Innovations for scalability

Synthetic reservoir application: Sequential EnIF



# The Synthetic case

- About 8.5 million parameters *p*.
  - About 72 static parameters
  - $4 \times 2D$  surfaces of size 123921
  - $9 \times 3D$  fields of size 886512
- Ensemble size n = 100.
- Number of responses m = 117, more if seismic is included.

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#### **Question:**

How to understand or inspect the update for a domain expert?



Joint assimilation is fast! But yields little understanding

- The engineer can know more than learned from data.
- KLD is *not* the objective of the engineer. Understanding, tuning, and a story?

#### Algorithm Sequential EnIF

- 1: Sample ensemble, estimate prior precision
- 2: for each batch  $d_b$  of observations do
- 3: Fit the sparse linear sub-map  $\hat{H}_b$
- 4: **for** each observation *k* in batch *b* **do**
- 5: Inspect  $\hat{H}_{b(k)}$ , tweak, understand effect, approve and a story
- 6: end for
- 7: Assimilate  $d_b$  using the (additive) EnIF update
- 8: end for

ea

Go to jupyter notebook



### If the demo did not work...





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• Statistial convergence!



- Statistial convergence!
- Structure and regularisation!



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- Statistial convergence!
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- Care about users: Sequential assimilation for understanding






- Akaike, Hirotugu (1974). "A new look at the statistical model identification". In: *IEEE transactions on automatic control* 19.6, pp. 716–723.
- Amestoy, Patrick R, Timothy A Davis, and Iain S Duff (2004). "Algorithm 837: AMD, an approximate minimum degree ordering algorithm". In: ACM Transactions on Mathematical Software (TOMS) 30.3, pp. 381–388.
- Claeskens, Gerda and Nils Lid Hjort (2008). "Model selection and model averaging". In: *Cambridge books*.
- Hastie, Trevor, Jonathan Taylor, et al. (2007). "Forward stagewise regression and the monotone lasso". In.
- Hastie, Trevor, Robert Tibshirani, et al. (2009). The elements of statistical learning: data mining, inference, and prediction. Vol. 2. Springer.



- Lindgren, Finn, Håvard Rue, and Johan Lindström (2011). "An explicit link between Gaussian fields and Gaussian Markov random fields: the stochastic partial differential equation approach". In: *Journal of the Royal Statistical Society: Series B (Statistical Methodology)* 73.4, pp. 423–498.
- Moore, John Barratt and B Anderson (1979). Optimal filtering. Prentice-Hall New York.
- Rozanov, Ju A (1977). "Markov random fields and stochastic partial differential equations". In: *Mathematics of the USSR-Sbornik* 32.4, p. 515.
- Rue, Havard and Leonhard Held (2005). Gaussian Markov random fields: theory and applications. Chapman and Hall/CRC.



- Takeuchi, Kei (1976). "Distribution of Information Statistics and Validity Criteria of Models". In: *Mathematical Science* 153, pp. 12–18.
- Tibshirani, Robert (1996). "Regression shrinkage and selection via the lasso". In: Journal of the Royal Statistical Society Series B: Statistical Methodology 58.1, pp. 267–288.

